Statistical Learning Theory for Modern Machine Learning

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Generalisation is the ability to 'perform' well on unseen data.

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- PAC: probably approximately correct [59]
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 ^m[large error] ≤ δ
 δ
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- PAC: probably approximately correct [59] Use a 'confidence parameter' δ : \mathbb{P}^m [large error] $\leq \delta$ δ is the probability of being misled by the training set
- Hence high confidence: \mathbb{P}^m [approximately correct] $\ge 1 \delta$

Error distribution picture



Learning algorithm $A : \mathcal{Z}^m \to \mathcal{H}$

• $\mathcal{Z} = \mathcal{X} \times \mathcal{Y}$ \mathcal{X} = set of inputs \mathcal{Y} = set of outputs (e.g. labels)

H = hypothesis class
 = set of predictors
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b these can be relaxed (but not in this talk)

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Actually these two goals interact with each other!

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Theoretical risk: (out-of-sample)

$$R_{\mathrm{out}}(h) = \mathbb{E}[\ell(h(X), Y)]$$

Examples:

- $\ell(h(X), Y) = \mathbf{1}[h(X) \neq Y]$: 0-1 loss (classification)
- $\ell(h(X), Y) = (Y h(X))^2$: square loss (regression)
- $\ell(h(X), Y) = (1 Yh(X))_+$: hinge loss
- $\ell(h(X), Y) = -\log(h(X))$: log loss (density estimation) TODO

■ Single hypothesis *h* (building block):

with probability $\ge 1 - \delta$, $R_{\text{out}}(h) \le R_{\text{in}}(h) + \sqrt{\frac{1}{2m} \log(\frac{1}{\delta})}$.

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 \longrightarrow Extension: PAC-Bayes allows to consider *distributions* over hypotheses.
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The risk measures $R_{in}(h)$ and $R_{out}(h)$ are extended by averaging: $R_{in}(Q) \equiv \int_{\mathcal{H}} R_{in}(h) \, dQ(h) \qquad R_{out}(Q) \equiv \int_{\mathcal{H}} R_{out}(h) \, dQ(h)$

 $\operatorname{KL}(\boldsymbol{Q} \| \boldsymbol{P}) = \underset{h \sim \boldsymbol{Q}}{\mathbf{E}} \ln \frac{\boldsymbol{Q}(h)}{\boldsymbol{P}(h)}$ is the Kullback-Leibler divergence.

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Prior	Bayesian inference Unique Statistical modelling (likelihood)	Posterior
Any distribution	PAC-Bayes Model-free	Any distribution
not depending on data	Inspired by the Bayesian update principle - Only depends on loss	(possibly) depending on data

"Prior": exploration mechanism of \mathcal{H} "Posterior" is the twisted prior after confronting the data

Prior

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- · PAC-Bayes: bounds hold for any distribution
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Data distribution

- PAC-Bayes: bounds hold for any distribution
- · Bayes: randomness lies in the noise model generating the output

A General PAC-Bayesian Theorem

 Δ -function: "distance" between $R_{in}(Q)$ and $R_{out}(Q)$

Convex function $\Delta : [0, 1] \times [0, 1] \rightarrow \mathbb{R}$.

General theorem

(Bégin et al. [7, 8], Germain [21])

For any distribution D on $\mathfrak{X} \times \mathfrak{Y}$, for any set \mathfrak{H} of voters, for any distribution P on \mathfrak{H} , for any $\delta \in (0, 1]$, and for any Δ -function, we have, with probability at least $1-\delta$ over the choice of $S \sim D^m$,

$$\forall Q \text{ on } \mathcal{H}: \quad \Delta\Big(R_{\mathrm{in}}(Q), R_{\mathrm{out}}(Q)\Big) \leqslant \frac{1}{m} \Big[\mathrm{KL}(Q \| P) + \ln \frac{\mathtt{J}_{\Delta}(m)}{\delta}\Big],$$

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where

$$\mathcal{J}_{\Delta}(m) = \sup_{r \in [0,1]} \left[\sum_{k=0}^{m} \underbrace{\binom{m}{k} r^{k} (1-r)^{m-k}}_{\operatorname{Bin}(k;m,r)} e^{m\Delta(\frac{k}{m},r)} \right]$$

Proof of the general theorem

General theorem

$$\Pr_{\mathcal{S}\sim D^m}\left(\forall \, \mathcal{Q} \text{ on } \mathcal{H}: \, \Delta\left(R_{\mathrm{in}}(\mathcal{Q}), R_{\mathrm{out}}(\mathcal{Q})\right) \leq \frac{1}{m}\left[\mathrm{KL}(\mathcal{Q}\|\mathcal{P}) + \ln \frac{\mathtt{J}_{\Delta}(m)}{\delta}\right]\right) \, \geq \, 1-\delta \, .$$

Proof ideas.

Change of Measure Inequality

For any P and Q on $\mathcal H,$ and for any measurable function $\varphi:\mathcal H\to\mathbb R,$ we have

$$-\ln\left(\mathop{\mathbf{E}}_{h\sim P} e^{\phi(h)}\right) = -\ln\mathop{\mathbf{E}}_{h\sim Q}\left(\frac{P(h)}{Q(h)}e^{\phi(h)}\right)$$
$$\leqslant \mathop{\mathbf{E}}_{h\sim Q}\ln\left(\frac{Q(h)}{P(h)}\right) - \mathop{\mathbf{E}}_{h\sim Q}\phi(h)$$
$$= \operatorname{KL}(Q||P) - \mathop{\mathbf{E}}_{h\sim Q}\phi(h).$$

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Markov's inequality

for a random variable X satisfying $X \ge 0$ $\Pr(X \ge a) \le \frac{\mathbb{E}X}{a} \iff \Pr(X \le \frac{\mathbb{E}X}{\delta}) \ge 1 - \delta$.

Proof of the general theorem

Probability of observing k misclassifications among m examples Given a voter h, consider a **binomial variable** of m trials with **success** $R_{out}(h)$:

$$\Pr_{S\sim D^m}\left(R_{\rm in}(h) = \frac{k}{m}\right) = \binom{m}{k} \left(R_{\rm out}(h)\right)^k \left(1 - R_{\rm out}(h)\right)^{m-k} = \operatorname{Bin}\left(k; m, R_{\rm out}(h)\right)$$

$$m \cdot \Delta \Big(\underset{h \sim Q}{\mathsf{E}} R_{\mathrm{in}}(h), \underset{h \sim Q}{\mathsf{E}} R_{\mathrm{out}}(h) \Big)$$

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Jensen's Inequality

 \leq

$$m \cdot \Delta \left(\underset{h \sim Q}{\mathsf{E}} R_{\mathrm{in}}(h), \underset{h \sim Q}{\mathsf{E}} R_{\mathrm{out}}(h) \right)$$
Jensen's Inequality
$$\leq \qquad \underset{h \sim Q}{\mathsf{E}} m \cdot \Delta \left(R_{\mathrm{in}}(h), R_{\mathrm{out}}(h) \right)$$
Change of measure
$$\leq \qquad \mathrm{KL}(Q \| P) + \ln \underset{h \sim P}{\mathsf{E}} e^{m\Delta \left(R_{\mathrm{in}}(h), R_{\mathrm{out}}(h) \right)}$$

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$$\Pr_{\mathcal{S}\sim D^m}\left(\forall Q \text{ on } \mathcal{H}: \Delta\left(R_{\text{in}}(Q), R_{\text{out}}(Q)\right) \leq \frac{1}{m}\left[\text{KL}(Q||P) + \ln \frac{J_{\Delta}(m)}{\delta}\right]\right) \geq 1-\delta.$$

Corollary

[...] with probability at least $1-\delta$ over the choice of $S \sim D^m$, for all Q on \mathcal{H} :

(a) $\operatorname{kl}\left(R_{\operatorname{in}}(Q), R_{\operatorname{out}}(Q)\right) \leq \frac{1}{m}\left[\operatorname{KL}(Q||P) + \ln \frac{2\sqrt{m}}{\delta}\right], \quad Langford and Seeger [31]$

$$\frac{\mathrm{kl}(q,\rho)}{=} \quad q \ln \frac{q}{\rho} + (1-q) \ln \frac{1-q}{1-\rho}$$

$$\Pr_{S\sim D^m}\left(\forall Q \text{ on } \mathcal{H}: \Delta\left(R_{\mathrm{in}}(Q), R_{\mathrm{out}}(Q)\right) \leq \frac{1}{m}\left[\mathrm{KL}(Q\|P) + \ln \frac{\mathtt{J}_{\Delta}(m)}{\delta}\right]\right) \geq 1-\delta.$$

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Catoni [11]

$$\begin{split} \mathrm{kl}(q,p) & \stackrel{\mathrm{def}}{=} \quad q \ln \frac{q}{p} + (1-q) \ln \frac{1-q}{1-p} \geqslant \ \mathbf{2}(q-p)^2 \,, \\ \Delta_c(q,p) & \stackrel{\mathrm{def}}{=} \quad -\ln[1-(1-e^{-c})\cdot p] - c \cdot q \,, \end{split}$$

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 $\begin{array}{ll} [\dots] \text{ with probability at least } 1-\delta \text{ over the choice of } S \sim D^m, \text{ for all } Q \text{ on } \mathcal{H}: \\ \hline \blacksquare & \operatorname{kl}\Big(R_{\operatorname{in}}(Q), R_{\operatorname{out}}(Q)\Big) \leq \frac{1}{m} \left[\operatorname{KL}(Q \| P) + \ln \frac{2\sqrt{m}}{\delta}\right], & \text{Langford and Seeger [31]} \\ \hline \blacksquare & R_{\operatorname{out}}(Q) \leq R_{\operatorname{in}}(Q)) + \sqrt{\frac{1}{2m} \left[\operatorname{KL}(Q \| P) + \ln \frac{2\sqrt{m}}{\delta}\right]}, & \text{McAllester [40, 43]} \\ \hline \blacksquare & R_{\operatorname{out}}(Q) \leq \frac{1}{1-e^{-c}} \left(c \cdot R_{\operatorname{in}}(Q) + \frac{1}{m} \left[\operatorname{KL}(Q \| P) + \ln \frac{1}{\delta}\right]\right), & \text{Catoni [11]} \\ \hline \blacksquare & R_{\operatorname{out}}(Q) \leq R_{\operatorname{in}}(Q) + \frac{1}{\lambda} \left[\operatorname{KL}(Q \| P) + \ln \frac{1}{\delta} + f(\lambda, m)\right]. & \text{Alquier et al. [4]} \end{array}$

$$\begin{split} & \mathrm{kl}(q,p) & \stackrel{\mathrm{def}}{=} & q \ln \frac{q}{p} + (1-q) \ln \frac{1-q}{1-p} \geqslant 2(q-p)^2 \,, \\ & \Delta_c(q,p) & \stackrel{\mathrm{def}}{=} & -\ln[1-(1-e^{-c}) \cdot p] - c \cdot q \,, \\ & \Delta_\lambda(q,p) & \stackrel{\mathrm{def}}{=} & \frac{\lambda}{m}(p-q) \,. \end{split}$$

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$$&\leqslant 2\sqrt{m}. \end{split}$$

■ Note that, in Line (1) of the proof, $\Pr_{S \sim D^m} (R_S(h) = \frac{k}{m})$ is replaced by the probability mass function of the binomial.

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- This is **only true if** the examples of *S* are drawn iid. (i.e., $S \sim D^m$)
- So this result is no longer valid in the non iid case, even if General Theorem is.
Linear classifiers

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- The specification of the centre for the posterior Q(w, μ) will be by a unit vector w and a scale factor μ.









Linear classifiers performance may be bounded by

$$\mathsf{KL}(\hat{Q}_{\mathcal{S}}(\mathbf{w},\mu) \| \mathbf{Q}_{\mathcal{D}}(\mathbf{w},\mu)) \leqslant \frac{\mathsf{KL}(P \| Q(\mathbf{w},\mu)) + \ln \frac{m+1}{\delta}}{m}$$

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■ SVM is deterministic classifier that exactly corresponds to sgn (E_{c~Q(w,µ)}[c(x)]) as centre of the Gaussian gives the same classification as halfspace with more weight.

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- SVM is deterministic classifier that exactly corresponds to sgn (E_{c~Q(w,µ)}[c(x)]) as centre of the Gaussian gives the same classification as halfspace with more weight.
- Hence its error bounded by 2Q_D(w, μ), since as observed above if x misclassified at least half of c ~ Q err.

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- $\blacksquare \operatorname{KL}(P \| Q) = \mu^2/2$

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- δ is the confidence
- The bound holds with probability 1 δ over the random i.i.d. selection of the training data.

Form of the SVM bound

Note that bound holds for all posterior distributions so that we can choose μ to optimise the bound

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- Note that bound holds for all posterior distributions so that we can choose μ to optimise the bound
- If we define the inverse of the KL by

 $\mathrm{KL}^{-1}(q, A) = \max\{p : \mathrm{KL}(q \| p) \leqslant A\}$

then have with probability at least $1-\delta$

$$\Pr\left(\langle \mathbf{w}, \boldsymbol{\varphi}(\mathbf{x}) \rangle \neq y\right) \leqslant 2\min_{\mu} \mathrm{KL}^{-1}\left(\mathbb{E}_m[\tilde{F}(\mu\gamma(\mathbf{x}, y))], \frac{\mu^2/2 + \ln\frac{m+1}{\delta}}{m}\right)$$

Gives SVM Optimisation

Primal form:

$$\min_{\mathbf{w},\xi_i} \begin{bmatrix} \frac{1}{2} \|\mathbf{w}\|^2 + C \sum_{i=1}^m \xi_i \end{bmatrix}$$

s.t. $y_i \mathbf{w}^T \phi(\mathbf{x}_i) \ge 1 - \xi_i$ $i = 1, \dots, m$
 $\xi_i \ge 0$ $i = 1, \dots, m$

Dual form:

$$\max_{\alpha} \left[\sum_{i=1}^{m} \alpha_{i} - \frac{1}{2} \sum_{i,j=1}^{m} \alpha_{i} \alpha_{j} y_{i} y_{j} \kappa(\mathbf{x}_{i}, \mathbf{x}_{j}) \right]$$

s.t. $0 \leq \alpha_{i} \leq C$ $i = 1, ..., m$

where $\kappa(\mathbf{x}_i, \mathbf{x}_j) = \langle \phi(\mathbf{x}_i), \phi(\mathbf{x}_j) \rangle$ and $\langle \mathbf{w}, \phi(\mathbf{x}) \rangle = \sum_{i=1}^m \alpha_i y_i \kappa(\mathbf{x}_i, \mathbf{x})$.



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 - defining the prior in terms of the data generating distribution (aka localised PAC-Bayes).

Learning the prior (1/3)

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- Introduce the learnt prior in the bound
- Compute stochastic error with remaining data







- Solve SVM with subset of patterns
- Prior in the direction w_r
- Posterior like PAC-Bayes Bound



- Solve SVM with subset of patterns
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- Posterior like PAC-Bayes Bound
- **New bound** depends on KL(P||Q)

SVM performance may be tightly bounded by

$$\mathsf{KL}(\hat{Q}_{\mathcal{S}}(\mathbf{w},\mu) \| \mathbf{Q}_{\mathcal{D}}(\mathbf{w},\mu)) \leq \frac{0.5 \|\mu \mathbf{w} - \eta \mathbf{w}_r\|^2 + \ln \frac{(m-r+1)J}{\delta}}{m-r}$$

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 $\hat{Q}_{S}(\mathbf{w}, \mu)$ stochastic measure of the training error on remaining data

 $\hat{Q}(\mathbf{w}, \mu)_{\mathcal{S}} = \mathbb{E}_{\mathbf{m}-\mathbf{r}}[\tilde{F}(\mu\gamma(\mathbf{x}, y))]$

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■ $0.5 \|\mu \mathbf{w} - \eta \mathbf{w}_r\|^2$ distance between prior and posterior

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Penalty term only dependent on the remaining data m - r

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s.t.
$$y_i \mathbf{w}^T \phi(\mathbf{x}_i) \ge 1 - \xi_i \qquad i = 1, \dots, m-r$$
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The p-SVM is only solved with the remaining points

 Determine the prior with a subset of the training examples to obtain w_r

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- **3 Margin** for the stochastic classifier \hat{Q}_s

$$\gamma(\mathbf{x}_j, y_j) = \frac{y_j \mathbf{w}^T \Phi(\mathbf{x}_j)}{\|\Phi(\mathbf{x}_j)\| \|\mathbf{w}\|} \qquad j = 1, \dots, m - r$$

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Linear search to obtain the optimal value of μ. This introduces an insignificant extra penalty term

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Bound for $\eta\text{-prior-SVM}$

- Prior is elongated along the line of w_r but spherical with variance 1 in other directions
- Posterior again on the line of w at a distance µ chosen to optimise the bound.
- Resulting bound depends on a benign parameter τ determining the variance in the direction w_r

$$\begin{split} \mathsf{KL}(\hat{Q}_{\mathcal{S}\backslash \mathcal{R}}(\mathbf{w},\mu) \| \mathcal{Q}_{\mathcal{D}}(\mathbf{w},\mu)) \leqslant \\ \frac{0.5(\ln(\tau^2) + \tau^{-2} - 1 + \mathcal{P}_{\mathbf{w}_r}^{\parallel}(\mu\mathbf{w} - \mathbf{w}_r)^2 / \tau^2 + \mathcal{P}_{\mathbf{w}_r}^{\perp}(\mu\mathbf{w})^2) + \ln(\frac{m-r+1}{\delta})}{\delta} \end{split}$$

m-r

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subject to

$$y_i(\mathbf{v} + \eta \mathbf{w}_r)^T \Phi(\mathbf{x}_i) \ge 1 - \xi_i$$
 $i = 1, \dots, m - r$
 $\xi_i \ge 0$ $i = 1, \dots, m - r$

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Results

		Classifier					
		SVM				ηPrior SVM	
Problem		2FCV	10FCV	PAC	PrPAC	PrPAC	τ-PrPAC
digits	Bound	_	_	0.175	0.107	0.050	0.047
	TE	0.007	0.007	0.007	0.014	0.010	0.009
waveform	Bound	_	_	0.203	0.185	0.178	0.176
	TE	0.090	0.086	0.084	0.088	0.087	0.086
pima	Bound	_	_	0.424	0.420	0.428	0.416
	TE	0.244	0.245	0.229	0.229	0.233	0.233
ringnorm	Bound	_	_	0.203	0.110	0.053	0.050
	TE	0.016	0.016	0.018	0.018	0.016	0.016
spam	Bound	_	_	0.254	0.198	0.186	0.178
	TE	0.066	0.063	0.067	0.077	0.070	0.072
Average	TE	0.0846	0.0834	0.081	0.0852	0.0832	0.0832

Take home messages

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- Model selection from the bounds is as good as 10FCV: in fact all but one of the PAC-Bayes model selections give better averages for TE.
- The better bounds do not appear to give better model selection best model selection is from the simplest bound.
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■ Consider *P* and *Q* are Gibbs-Boltzmann distributions

$$P(h) := \frac{1}{Z'} e^{-\gamma \operatorname{risk}(h)} \qquad Q(h) := \frac{1}{Z} e^{-\gamma \operatorname{risk}_{\mathcal{S}}(h)}$$

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These distributions are hard to work with since we cannot apply the bound to a single weight vector, but the bounds can be very tight:

$$\mathcal{K}L_{+}(\hat{Q}_{\mathcal{S}}(\gamma) \| \mathcal{Q}_{\mathcal{D}}(\gamma)) \leqslant \frac{1}{m} \left(\frac{\gamma}{\sqrt{m}} \sqrt{\ln \frac{8\sqrt{m}}{\delta}} + \frac{\gamma^{2}}{4m} + \ln \frac{4\sqrt{m}}{\delta} \right)$$

with the only uncertainty the dependence on γ .

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 - Trick here is that the error measures only depend on the posterior Q, while the bound depends on KL between posterior and prior: an estimate of this KL is made without knowing the prior explicitly
- the Gibbs distributions are hard to sample from so not easy to work with this bound.

Other distribution defined priors

An alternative distribution defined prior for an SVM is to place symmetrical Gaussian at the weight vector:
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 w_ρ = E_{(x,y)~D}(yφ(x)) to give distributions that are easier to work with, but results not impressive...
- What if we were to take the expected weight vector returned from a random training set of size *m*: then the KL between posterior and prior is related to the concentration of weight vectors from different training sets
- This is connected to stability...

Outline



Stability

Uniform hypothesis sensitivity β at sample size *m*:

$$\| A(z_{1:m}) - A(z'_{1:m}) \| \leq \beta \sum_{i=1}^{m} \mathbf{1}[z_i \neq z'_i]$$

$$\begin{array}{ll} (z_1,\ldots,z_m) & (z_1',\ldots,z_m') \\ \blacksquare & \mathsf{A}(z_{1:m}) \in \mathcal{H} \text{ normed space} & \blacksquare \text{ Lipschitz} \\ \blacksquare & w_m = \mathsf{A}(z_{1:m}) \text{ 'weight vector'} & \blacksquare \text{ smoothness} \end{array}$$

Uniform loss sensitivity β at sample size *m*:

 $|\ell(\mathsf{A}(z_{1:m}), z) - \ell(\mathsf{A}(z'_{1:m}), z)| \leq \beta \sum_{i=1}^{m} \mathbf{1}[z_i \neq z'_i]$

worst-case

distribution-insensitive

data-insensitive

Open: data-dependent?

If A has sensitivity β at sample size *m*, then for any $\delta \in (0, 1)$, w.p. $\ge 1 - \delta$, $R_{\text{out}}(h) \le R_{\text{in}}(h) + \varepsilon(\beta, m, \delta)$

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- the intuition is that if individual examples do not affect the loss of an algorithm then it will be concentrated
- can be applied to kernel methods where β is related to the regularisation constant, but bounds are quite weak
- question: algorithm output is highly concentrated stronger results?

Stability + PAC-Bayes I

If A has uniform hypothesis stability β at sample size *n*, then for any $\delta \in (0, 1)$, w.p. $\ge 1 - 2\delta$,

$$\mathrm{KL}(R_{\mathrm{in}}(Q) \| R_{\mathrm{out}}(Q)) \leqslant \frac{\frac{n\beta^2}{2\sigma^2} \left(1 + \sqrt{\frac{1}{2}\log\left(\frac{1}{\delta}\right)}\right)^2 + \log\left(\frac{n+1}{\delta}\right)}{n}$$

Gaussian randomization

•
$$P = \mathcal{N}(\mathbb{E}[W_n], \sigma^2 I)$$

• $Q = \mathcal{N}(W_n, \sigma^2 I)$ • $\mathrm{KL}(Q \| P) = \frac{1}{2\sigma^2} \| W_n - \mathbb{E}[W_n] \|^2$

Main proof components:

• w.p. $\geq 1 - \delta$, $\operatorname{KL}(R_{\operatorname{in}}(Q) || R_{\operatorname{out}}(Q)) \leq \frac{\operatorname{KL}(Q || Q_0) + \log\left(\frac{n+1}{\delta}\right)}{n}$

• w.p.
$$\geq 1 - \delta$$
, $\|W_n - \mathbb{E}[W_n]\| \leq \sqrt{n} \beta \left(1 + \sqrt{\frac{1}{2}\log(\frac{1}{\delta})}\right)$

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- Use second part of data to perform an optimisation of a PAC-Bayes bound
- Different ways to choose approximations to the KL term between empirical and true risk: the relaxed Pinsker inequality reads:

 $kl(\hat{p}||p) \ge 2(p-\hat{p})^2 \quad \text{for } \hat{p}, p \in (0, 1), \qquad (f_{classic}) \qquad (2)$

while the refined Pinsker inequality takes the form:

$$\mathrm{kl}(\hat{\boldsymbol{\rho}} \| \boldsymbol{\rho}) \geqslant \frac{(\boldsymbol{\rho} - \hat{\boldsymbol{\rho}})^2}{2\boldsymbol{\rho}} \quad \text{ for } \hat{\boldsymbol{\rho}}, \boldsymbol{\rho} \in (0, 1), \ \hat{\boldsymbol{\rho}} < \boldsymbol{\rho}. \quad (f_{\mathrm{quad}}) \quad (3)$$

■ f_{λ} based on the λ bound and f_{bbb} based on variational inference.

Model Selection Results



Figure: Model selection results from more than 600 runs with different hyper-parameters. The architecture used is a CNN with Gaussian data-dependent priors. We use a reduced subset of MNIST for these experiments (10% of training data).
Training and Generalisation Results

Setup		Risk cert.		Stch. pred.		Det. pred.		Ens. pred.		Prior	
Arch.	Prior	Obj.	ℓ ^{x-e}	ℓ ⁰¹	х-е	01 err.	х-е	01 err.	х-е	01 err.	01 err.
FCN		f _{quad}	.2033	.3155	.0268	.0921	.0137	.0558	.0007	.0572	.8792
	Rand.Init.	flambda	.2326	.3275	.0211	.0732	.0077	.0429	.0004	.0448	.8792
	(Gaussian)	$f_{\rm classic}$.1749	.3304	.0407	.1411	.0204	.0851	.0009	.0868	.8792
		f _{bbb}	.5163	.5516	.0088	.0293	.0038	.0172	.0003	.0178	.8792
		f _{quad}	.0146	.0279	.0084	.0202	.0032	.0186	.0002	.0189	.0202
	Learnt	$f_{\rm lambda}$.0201	.0354	.0082	.0196	.0071	.0185	.0001	.0185	.0202
	(Gaussian)	$f_{\rm classic}$.0141	.0284	.0101	.0230	.0089	.0189	.0002	.0191	.0202
		f _{bbb}	.0788	.0968	.0063	.0179	.0066	.0153	.0001	.0153	.0202
	-	<i>f</i> erm	-	-	-	-	.0101	.0152	-	-	-
CNN		f _{quad}	.1453	.2165	.0143	.0513	.0062	.0257	.0003	.0261	.9478
	Rand.Init.	$f_{\rm lambda}$.1583	.2202	.0109	.0397	.0056	.0207	.0003	.0211	.9478
	(Gaussian)	$f_{\rm classic}$.1260	.2277	.0253	.0869	.0111	.0425	.0006	.0421	.9478
		f _{bbb}	.3400	.3645	.0039	.0154	.0016	.0088	.0001	.0092	.9478
		f _{quad}	.0078	.0155	.0045	.0104	.0003	.0105	.0001	.0104	.0104
	Learnt	$f_{\rm lambda}$.0095	.0186	.0044	.0106	.0047	.0098	.0000	.0100	.0104
	(Gaussian)	$f_{\rm classic}$.0083	.0166	.0049	.0123	.0048	.0103	.0001	.0103	.0104
		f _{bbb}	.0447	.0538	.0040	.0104	.0043	.0082	.0002	.0082	.0104
	-	f _{erm}	-	-	-	-	.0081	.0092	-	-	-

Table: MNIST using Gaussian priors. The table includes two architectures (FCN and CNN), two priors (a data-free prior , and a data-dependent prior) and four training objectives.

Training and Generalisation Results

Setup			Risk cert.		Stch. pred.		Det. pred.		Ens. pred.		Prior
Arch.	Prior	Obj.	ℓ ^{x-e}	ℓ01	х-е	01 err.	x-e	01 err.	х-е	01 err.	01 err.
CNN (9 layers)	Learnt (50% data)	fquad	.1296	.3034	.0903	.2452	.0726	.2439	.0024	.2413	.2518
		t _{lambda}	.1742	.3730 .2901	.0689 .0931	.2307 .2537	.0609	.2225 .2437	.0018	.2133 .2332	.2518 .2518
		$f_{\rm classic}$ $f_{\rm bbb}$.8096	.8633	.0715	.2198	.0735	.2437	.0023	.2332	.2518
	Learnt (70% data)	f _{quad}	.1017	.2502	.0816	.2137	.0928	.2137	.0023	.2100	.2169
		$f_{ m lambda}$.1414	.3128	.0708	.2081	.0767	.2061	.0021	.2049	.2169
		$f_{\rm classic}$.0957	.2377	.0862	.2161	.0827	.2167	.0021	.2135	.2169
		f _{bbb}	.6142	.6965	.0708	.1979	.0562	.1992	.0019	.1944	.2169
	-	f _{erm}	-	-	-	-	.1400	.1946	-	-	-
	Learnt (50% data)	<i>f</i> _{quad}	.0867	.2174	.0584	.1668	.0538	.1662	.0014	.1653	.1688
		$f_{\rm lambda}$.1217	.2707	.0506	.1618	.0417	.1639	.0015	.1622	.1688
		$f_{\rm classic}$.0782	.1954	.0652	.1686	.0594	.1692	.0013	.1674	.1688
CNN		<i>f</i> _{bbb}	.6069	.7066	.0468	.1553	.0412	.1530	.0012	.1517	.1688
(15 layers)	Learnt (70% data)	<i>f</i> _{quad}	.0756	.1806	.0559	.1463	.0391	.1469	.0016	.1449	.1490
(13 layers)		$f_{\rm lambda}$.0922	.2121	.0500	.1437	.0507	.1449	.0012	.1438	.1490
		$f_{\rm classic}$.0703	.1667	.0615	.1475	.0551	.1480	.0010	.1476	.1490
		f _{bbb}	.4481	.5572	.0455	.1413	.0395	.1405	.0008	.1409	.1490
	-	f _{erm}	-	-	-	-	.0957	.1413	-	-	-

Table: Train and test set results on CIFAR-10 using Gaussian priors, three deep CNN architectures and two percentages of data used to build the data-dependent prior (50% and 70%, i.e. 25.000 and 35.000 examples).

A flexible framework

A flexible framework

Since 1997, PAC-Bayes has been successfully used in many machine learning settings (this list is by no means exhaustive).

Statistical learning theory Audibert and Bousquet [6], Catoni [9, 10], Guedj [25], Guedj and Pujol [27], Maurer [39], McAllester [41, 42, 44, 45], Mhammedi et al. [46], Seeger [51, 52], Shawe-Taylor

and Williamson [56], Thiemann et al. [58]

- SVMs & linear classifiers Germain et al. [19], Langford and Shawe-Taylor [32], McAllester [44]
- Supervised learning algorithms reinterpreted as bound minimizers Ambroladze et al. [5], Germain et al. [22], Shawe-Taylor and Hardoon [57]
- High-dimensional regression Alquier and Biau [1], Alquier and Lounici [2], Guedj and Robbiano [24], Guedj and Alquier [26], Li et al. [35]
 Classification Catoni [9, 10], Lacasse et al. [30], Langford and Shawe-Taylor [32], Parrado-Hernández et al. [49]

A flexible framework

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Transductive learning, domain adaptation Bégin et al. [7]. Derbeko et al. [12], Germain et al. [20], Nozawa et al. [48] Non-iid or heavy-tailed data Alquier and Guedj [3], Holland [29], Lever et al. [34], Seldin et al. [54, 55] Density estimation Higgs and Shawe-Taylor [28], Seldin and Tishby [53] Reinforcement learning Fard and Pineau [16], Fard et al. [17], Ghavamzadeh et al. [23], Seldin et al. [54, 55] Sequential learning Gerchinovitz [18], Li et al. [36] Algorithmic stability, differential privacy Dziugaite and Roy [13, 14], London [37], London et al. [38], Rivasplata et al. [50] Deep neural networks Dziugaite and Roy [15], Letarte et al. [33], Neyshabur

et al. [47], Zhou et al. [60]

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- Modern machine learning appears to contradict many of the conclusions of statistical learning theory
- Modelling learning in a more refined way leads to bounds that overcome this contradiction and throw light on different ingredients in achieving good test performance
- Can drive algorithms to give improved bounds and state of the art performance
- Many other aspects of deep learning still remain to be captured by theoretical analysis

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